

**Problem 1 (Kang, 2001)**

Let  $X_1$  and  $X_2$  be independent random variables denoting the two picks that are uniformly distributed over the interval  $[0, a]$ . Let  $G(a) \equiv E[X^2] \equiv E[(\max(X_1, X_2))^2]$ . Suppose  $a < X_1 \leq a + \varepsilon$  and  $0 \leq X_2 \leq a$ .  $G(a + \varepsilon)$  for this case is computed as follows:

$$\begin{aligned} G(a + \varepsilon) &= E[(\max(X_1, X_2))^2] = E[X_1^2] = \int_a^{a+\varepsilon} x_1^2 f_{X_1}(x_1) dx_1 = \frac{1}{\varepsilon} \int_a^{a+\varepsilon} x_1^2 dx_1 \\ &= \frac{1}{\varepsilon} \left[ \frac{1}{3} x_1^3 \right]_a^{a+\varepsilon} = a^2 + a\varepsilon + o(\varepsilon), \end{aligned}$$

where  $o(\varepsilon)$  represents higher order terms of  $\varepsilon$  satisfying  $\lim_{\varepsilon \rightarrow 0} \frac{o(\varepsilon)}{\varepsilon} = 0$ . Ignoring  $o(\varepsilon)$ , we have the following table that summarizes  $G(a + \varepsilon)$ 's.

Case	Probability of a case	$G(a + \varepsilon)$ given a case
$0 \leq X_1 \leq a, 0 \leq X_2 \leq a$	$\frac{a}{a+\varepsilon} \cdot \frac{a}{a+\varepsilon} = \left(\frac{a}{a+\varepsilon}\right)^2$	$G(a)$
$a < X_1 \leq a + \varepsilon, 0 \leq X_2 \leq a$	$\frac{\varepsilon}{a+\varepsilon} \cdot \frac{a}{a+\varepsilon} = \frac{\varepsilon a}{(a+\varepsilon)^2}$	$a^2 + a\varepsilon$
$0 \leq X_1 \leq a, a < X_2 \leq a + \varepsilon$	$\frac{a}{a+\varepsilon} \cdot \frac{\varepsilon}{a+\varepsilon} = \frac{\varepsilon a}{(a+\varepsilon)^2}$	$a^2 + a\varepsilon$
$a < X_1 \leq a + \varepsilon, a < X_2 \leq a + \varepsilon$	$\frac{\varepsilon}{a+\varepsilon} \cdot \frac{\varepsilon}{a+\varepsilon} = \left(\frac{\varepsilon}{a+\varepsilon}\right)^2$	We do not care.

Using the total expectation theorem, we obtain

$$\begin{aligned} G(a + \varepsilon) &= G(a) \left(\frac{a}{a + \varepsilon}\right)^2 + (a^2 + a\varepsilon) \frac{\varepsilon a}{(a + \varepsilon)^2} + (a^2 + a\varepsilon) \frac{\varepsilon a}{(a + \varepsilon)^2} + o(\varepsilon^2) \\ &\approx G(a) \left(\frac{a}{a + \varepsilon}\right)^2 + 2(a^2 + a\varepsilon) \frac{\varepsilon a}{(a + \varepsilon)^2} \end{aligned}$$

From the formula of the sum of an infinite geometric series, we know

$$\frac{a}{a + \varepsilon} = \frac{1}{1 + \frac{\varepsilon}{a}} = 1 - \frac{\varepsilon}{a} + \left(\frac{\varepsilon}{a}\right)^2 - \left(\frac{\varepsilon}{a}\right)^3 + \dots$$

Ignoring higher order terms of  $\varepsilon$ , we have

$$\frac{a}{a + \varepsilon} \approx 1 - \frac{\varepsilon}{a}$$

This gives the following approximations:

$$\left(\frac{a}{a+\varepsilon}\right)^2 \approx \left(1 - \frac{\varepsilon}{a}\right)^2 = 1 - \frac{2\varepsilon}{a} + \frac{\varepsilon^2}{a^2} \approx 1 - \frac{2\varepsilon}{a}$$

$$\frac{\varepsilon a}{(a+\varepsilon)^2} = \frac{\varepsilon}{a} \left(\frac{a}{a+\varepsilon}\right)^2 \approx \frac{\varepsilon}{a} \left(1 - \frac{2\varepsilon}{a}\right) = \frac{\varepsilon}{a} - \frac{2\varepsilon^2}{a^2} \approx \frac{\varepsilon}{a}$$

Therefore, we can rewrite  $G(a+\varepsilon)$  as

$$G(a+\varepsilon) \approx G(a) \left(1 - \frac{2\varepsilon}{a}\right) + 2(a^2 + a\varepsilon) \cdot \frac{\varepsilon}{a} \approx G(a) \left(1 - \frac{2\varepsilon}{a}\right) + 2a\varepsilon$$

Rearranging terms, we have

$$\frac{G(a+\varepsilon) - G(a)}{\varepsilon} = -\frac{2G(a)}{a} + 2a$$

If  $\varepsilon \rightarrow 0$ , we have the following differential equation:

$$G'(a) = -\frac{2G(a)}{a} + 2a$$

Let  $G(a) = Aa^2 + Ba + C$ . Since  $G(0) = 0$ , we have  $C = 0$ . From the differential equation,

$$2Aa + B = \frac{-2Aa^2 - 2Ba}{a} + 2a = (2 - 2A)a - 2B$$

It gives  $A = \frac{1}{2}$  and  $B = 0$ . Therefore

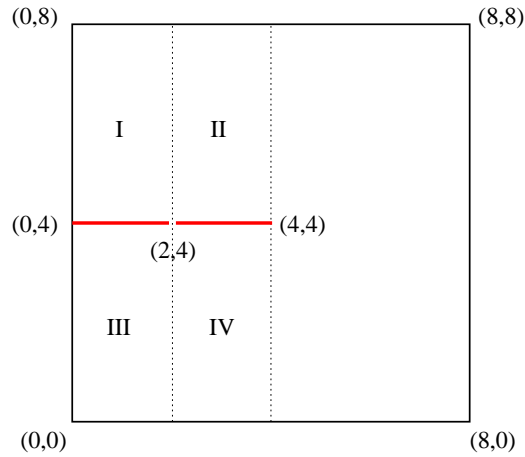
$$G(a) \equiv E[X^2] \equiv E[(\max(X_1, X_2))^2] = \frac{a^2}{2}$$

## Problem 2 (Kang, 2001)

Let  $X_1$  and  $X_2$  denote the locations of the response vehicle and an event, respectively.

- (i) The probability that the presence of the barrier increases the grid distance the vehicle must travel to the event,  $P(B)$ , is given by (refer to the figure below)

$$\begin{aligned} P(B) &= P(X_1 \in \text{I}, X_2 \in \text{III}) + P(X_1 \in \text{III}, X_2 \in \text{I}) + P(X_1 \in \text{II}, X_2 \in \text{IV}) + P(X_1 \in \text{IV}, X_2 \in \text{II}) \\ &= 2P(X_1 \in \text{I}, X_2 \in \text{III}) + 2P(X_1 \in \text{II}, X_2 \in \text{IV}) \\ &= 2 \cdot \frac{8}{64} \cdot \frac{8}{64} + 2 \cdot \frac{8}{64} \cdot \frac{8}{64} = \frac{1}{16} \end{aligned}$$



(ii) Let  $D$  be the travel distance without the barrier. We know from class

$$E[D] = E[D_x] + E[D_y] = \frac{1}{3} \times 8 + \frac{1}{3} \times 8 = \frac{16}{3}$$

Let  $D^e$  denote the extra distance the vehicle should travel due to the barrier. Let us first compute  $E[D^e \mid X_1 \in \text{I}, X_2 \in \text{III}]$ . There is no extra travel distance in the  $y$  axis, i.e.  $E[D_y^e \mid X_1 \in \text{I}, X_2 \in \text{III}] = 0$ . We also know from class that the extra travel distance in the  $x$  axis,  $E[D_x^e \mid X_1 \in \text{I}, X_2 \in \text{III}]$ , is  $\frac{2}{3} \times 2 = \frac{4}{3}$ . Hence

$$E[D^e \mid X_1 \in \text{I}, X_2 \in \text{III}] = \frac{4}{3}$$

By symmetry,

$$E[D^e \mid X_1 \in \text{III}, X_2 \in \text{I}] = \frac{4}{3}$$

Now consider  $E[D^e \mid X_1 \in \text{II}, X_2 \in \text{IV}]$ . As before,  $E[D_y^e \mid X_1 \in \text{II}, X_2 \in \text{IV}] = 0$ . To compute  $E[D_x^e \mid X_1 \in \text{II}, X_2 \in \text{IV}]$ , we should note that it is possible to travel through the both ends of the barrier spanning from  $(2, 4)$  to  $(4, 4)$ . We saw in a problem set when travel is allowed through the both ends of the barrier, the extra travel distance is  $\frac{1}{3}$  times the length of the barrier (refer to Problem 3.14 in the textbook). Therefore,

$$E[D^e \mid X_1 \in \text{II}, X_2 \in \text{IV}] = \frac{1}{3} \times 2 = \frac{2}{3}$$

By symmetry,

$$E[D^e \mid X_1 \in \text{IV}, X_2 \in \text{II}] = \frac{1}{3} \times 2 = \frac{2}{3}$$

$E[D^e]$  is then computed by

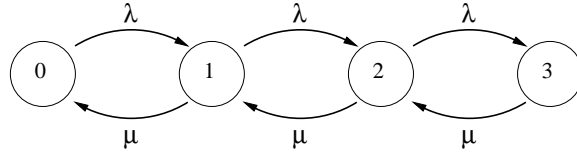
$$\begin{aligned}
E[D^e] &= E[D^e \mid X_1 \in \text{I}, X_2 \in \text{III}] P(X_1 \in \text{I}, X_2 \in \text{III}) + \\
&\quad E[D^e \mid X_1 \in \text{III}, X_2 \in \text{I}] P(X_1 \in \text{III}, X_2 \in \text{I}) + \\
&\quad E[D^e \mid X_1 \in \text{II}, X_2 \in \text{IV}] P(X_1 \in \text{II}, X_2 \in \text{IV}) + \\
&\quad E[D^e \mid X_1 \in \text{IV}, X_2 \in \text{II}] P(X_1 \in \text{IV}, X_2 \in \text{II}) \\
&= \frac{4}{3} \cdot \frac{8}{64} \cdot \frac{8}{64} + \frac{4}{3} \cdot \frac{8}{64} \cdot \frac{8}{64} + \frac{2}{3} \cdot \frac{8}{64} \cdot \frac{8}{64} + \frac{2}{3} \cdot \frac{8}{64} \cdot \frac{8}{64} = \frac{1}{16}
\end{aligned}$$

The expected total travel distance,  $E[D']$ , is therefore given by

$$E[D'] = E[D] + E[D^e] = \frac{16}{3} + \frac{1}{16} = \frac{259}{48}$$

**Problem 3 (Odoni, 2001)**

- (a) If the PDF of service time is negative exponential, the state transition diagram of Vincent's barbershop queueing system is given by



For the case where  $\lambda = \mu$ , we have the following balance equations and normalization equation:

$$\begin{aligned}
P_0 &= P_1 \\
P_1 &= P_2 \\
P_2 &= P_3 \\
P_0 + P_1 + P_2 + P_3 &= 1
\end{aligned}$$

Solving equations, we have  $P_0 = P_1 = P_2 = P_3 = \frac{1}{4}$ . The expected number of customers in the barbershop is

$$L = 1 \times P_1 + 2 \times P_2 + 3 \times P_3 = \frac{6}{4} = 1.5$$

- (b) Suppose there are  $k$  chairs (including the barber's chair) in the barbershop, which is to be determined. The balance equations and the normalization equation in this case are given by

$$P_n = P_{n+1}, \quad \text{for } n = 0, 1, \dots, k-1$$

$$P_0 + P_1 + \dots + P_k = 1$$

Clearly,  $P_0 = P_1 = \dots = P_k = \frac{1}{k+1}$ . To make sure that at least 92% of his prospective customers become actual customers, the probability that a new customer finds all chairs occupied,  $P_k$ , should be less than 8%.

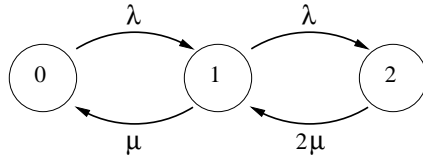
$$P_k = \frac{1}{k+1} < 0.08 \Rightarrow k > 11.5$$

The minimum number of chairs he will need in the shop is 12.

- (c) The state transition diagram for SIRO will be the same as that for FIFO. This means that the steady-state probabilities  $P_n$  will be identical in the two cases. Then  $L = \sum_{n=0}^3 nP_n$  will be the same in the two cases. Therefore,  $W = \frac{L}{\lambda} = \frac{L}{\lambda(1-P_3)}$  will be the same.

**Problem 4 (Odoni, 2001)**

- (a) The state transition diagram of this M/M/2 queueing system is



The balance equations and the normalization equation are

$$\lambda P_0 = \mu P_1$$

$$\lambda P_1 = 2\mu P_2$$

$$P_0 + P_1 + P_2 = 1$$

$P_1 = \frac{\lambda}{\mu} P_0 = \rho P_0$ .  $P_2 = \frac{\lambda}{2\mu} P_1 = \frac{1}{2} \rho P_1 = \frac{1}{2} \rho^2 P_0$ . Using the normalization equation,

$$P_0 + \rho P_0 + \frac{1}{2} \rho^2 P_0 = P_0 \left( 1 + \rho + \frac{1}{2} \rho^2 \right) = 1 \Rightarrow P_0 = \frac{1}{1 + \rho + \frac{1}{2} \rho^2}$$

The expected number of men who are busy serving a customer at any given time is given by

$$1 \times P_1 + 2 \times P_2 = \frac{\rho}{1 + \rho + \frac{1}{2} \rho^2} + \frac{\rho^2}{1 + \rho + \frac{1}{2} \rho^2} = \frac{\rho + \rho^2}{1 + \rho + \frac{1}{2} \rho^2}$$

(b) Using the data collected, we have the following equation:

$$\begin{aligned}\frac{\rho + \rho^2}{1 + \rho + \frac{1}{2}\rho^2} = \frac{8,000}{10,000} = 0.8 &\Rightarrow 0.8 + 0.8\rho + 0.4\rho^2 = \rho + \rho^2 \\ &\Rightarrow 0.6\rho^2 + 0.2\rho - 0.8 = 0 \\ &\Rightarrow \rho^2 + \frac{1}{3}\rho - \frac{4}{3} = 0\end{aligned}$$

It gives  $\rho = 1$  (the other root,  $-\frac{4}{3}$ , is meaningless). Note that the actual arrival rate of customers is  $\lambda' = \lambda(1 - P_2)$ . Since 40,000 customers received service during 10,000 hours,

$$\lambda(1 - P_2) = \frac{40,000}{10,000} = 4$$

Since  $\rho = 1$ , we have  $P_2 = \frac{\frac{1}{2}\rho^2}{1 + \rho + \frac{1}{2}\rho^2} = \frac{1}{5}$ . Therefore

$$\lambda = \frac{4}{(1 - P_2)} = \frac{4}{4/5} = 5$$

The number of customers lost during these 10,000 hours is

$$\lambda P_2 \times 10,000 = 5 \times \frac{1}{5} \times 10,000 = 10,000$$