

14.581 MIT PhD International Trade
—Lecture 14: Firm-Level Trade
(Empirics Part II)—

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Plan for Today's Lecture on Firm-Level Trade

1. Trade flows: intensive and extensive margins
2. Exporting across to multiple destinations

Intensive and Extensive Margins in Trade Flows

- With access to micro data on trade flows at the firm-level, a key question to ask is whether trade flows expand over time (or look bigger in the cross-section) along the:
 - Intensive margin: the same firms (or product-firms) from country i export more volume (and/or charge higher prices—we can also decompose the intensive margin into these two margins) to country j .
 - Extensive margin: new firms (or product-firms) from country i are penetrating the market in country j .
- This is really just a decomposition—we can and should expect trade to expand along both margins.
- Recently some papers have been able to look at this.
 - A rough lesson from these exercises is that the extensive margin seems more important (in a pure ‘accounting’ sense).

Bernard, Jensen, Redding and Schott (2007): Exporters

Data from US manufacturing firms. The coefficients in columns 2-4 sum (across columns) to those in column 1.

Table 6

Gravity and Aggregate U.S. Exports, 2000

| | <i>Log of total exports value</i> | <i>Log of number of exporting firms</i> | <i>Log of number of exported products</i> | <i>Log of export value per product per firm</i> |
|-----------------|-----------------------------------|---|---|---|
| Log of GDP | 0.98 (0.04) | 0.71 (0.04) | 0.52 (0.03) | -0.25 (0.04) |
| Log of distance | -1.36 (0.17) | -1.14 (0.16) | -1.06 (0.15) | 0.84 (0.19) |
| Observations | 175 | 175 | 175 | 175 |
| R ² | 0.82 | 0.74 | 0.64 | 0.25 |

Sources: Data are from the 2000 Linked-Longitudinal Firm Trade Transaction Database (LFTTD).

Notes: Each column reports the results of a country-level ordinary least squares regression of the dependent variable noted at the top of each column on the covariates noted in the first column. Results for the constant are suppressed. Standard errors are noted below each coefficient. Products are defined as ten-digit Harmonized System categories. All results are statistically significant at the 1 percent level.

Bernard, Jensen, Redding and Schott (2007): Importers

Data from US manufacturing firms. The coefficients in columns 2-4 sum (across columns) to those in column 1.

Table 9

Gravity and Aggregate U.S. Imports, 2000

| | <i>Log of total import value</i> | <i>Log of number of importing firms</i> | <i>Log of number of imported products</i> | <i>Log of import value per product per firm</i> |
|------------------------|--|---|---|---|
| <i>Log of GDP</i> | 1.14*** (0.06) | 0.82*** (0.03) | 0.71*** (0.03) | -0.39*** (0.05) |
| <i>Log of Distance</i> | -0.73*** (0.27) | -0.43*** (0.15) | -0.61*** (0.15) | 0.31 (0.24) |
| <i>Observations</i> | 175 | 175 | 175 | 175 |
| <i>R²</i> | 0.69 | 0.78 | 0.74 | 0.25 |

Sources: Data are from the 2000 Linked-Longitudinal Firm Trade Transaction Database (LFTTD).

Notes: Each column reports the results of a country-level ordinary least squares regression of the dependent variable noted at the top of each column on the covariates listed on the left. Results for constants are suppressed. Standard errors are noted below each coefficient. Products are defined as ten-digit Harmonized System categories.

*, **, and *** represent statistical significance at the 10, 5, and 1 percent levels, respectively.

Crozet and Koenig (CJE, 2010)

Data from French manufacturing firms trading internationally, by domestic region j .
 Extensive margin biased down by inclusion of only firms over 20 workers.

**Decomposition of French Aggregate Industrial Exports
 (34 Industries - 159 Countries - 1986 to 1992)**

| | All Firms >20 Employees | | Single-Region Firms >20 Employees | |
|------------------------|---|--|---|--|
| | (1) | (2) | (3) | (4) |
| | Average Shipment $\ln(M_{kjt} / N_{kjt})$ | Number of Shipments $\ln(N_{kjt})$ | Average Shipment $\ln(M_{kjt} / N_{kjt})$ | Number of Shipments $\ln(N_{kjt})$ |
| $\ln(\text{GDP}_{kj})$ | 0.461 ^a | 0.417 ^a | 0.421 ^a | 0.417 ^a |
| | (0.007) | (0.007) | (0.007) | (0.008) |
| $\ln(\text{Dist}_j)$ | -0.325 ^a | -0.446 ^a | -0.363 ^a | -0.475 ^a |
| | (0.013) | (0.009) | (0.012) | (0.009) |
| Contig_j | -0.064 ^c | -0.007 | 0.002 | 0.190 ^a |
| | (0.035) | (0.032) | (0.038) | (0.036) |
| Colony_j | 0.100 ^a | 0.466 ^a | 0.141 ^a | 0.442 ^a |
| | (0.032) | (0.025) | (0.035) | (0.027) |
| French_j | 0.213 ^a | 0.991 ^a | 0.188 ^a | 1.015 ^a |
| | (0.029) | (0.028) | (0.032) | (0.028) |
| N | 23553 | 23553 | 23553 | 23553 |
| R^2 | 0.480 | 0.591 | 0.396 | 0.569 |

Note: These are OLS estimates with year and industry dummies. Robust standard errors in parentheses with ^a, ^b and ^c denoting significance at the 1%, 5%, and 10% level respectively.

Hilberry and Hummels (EER, 2008)

Data on intra-national US commodity shipping (zipcode-to-zipcode, with firm identifiers).

Decomposing Spatial Frictions (5-digit zip code data)

| | Dist | Dist ² | Ownzip | Ownstate | Constant | Adj. R ² | N | ϵ_D |
|---|-------------------|-------------------|-------------------|-------------------|--------------------|---------------------|---------|--------------|
| Value (T_{ij}) | -0.137 (0.009) | -0.004 (0.001) | 1.102 (0.030) | -0.024 (0.007) | -13.393 (0.026) | 0.01 | 1290788 | -0.187 |
| # of shipments (N_{ij}) | -0.294 (0.002) | 0.017 (0.000) | 0.883 (0.008) | 0.043 (0.002) | -1.413 (0.007) | 0.10 | 1290840 | -0.081 |
| # of trading pairs (N_{ij}^f) | -0.159 (0.002) | 0.008 (0.000) | 0.540 (0.007) | 0.029 (0.002) | -0.888 (0.006) | 0.05 | 1290840 | -0.059 |
| # of commodities (N_{ij}^c) | -0.135 (0.001) | 0.009 (0.000) | 0.342 (0.003) | 0.014 (0.001) | -0.525 (0.003) | 0.10 | 1290840 | -0.022 |
| Avg. Value ($\bar{P}\bar{Q}_{ij}$) | 0.157 (0.008) | -0.021 (0.001) | 0.219 (0.028) | -0.067 (0.006) | -11.980 (0.024) | 0.00 | 1290788 | -0.106 |
| avg. price (\bar{P}_{ij}) | -0.032 (0.007) | 0.036 (0.001) | -0.115 (0.024) | -0.154 (0.006) | 0.021 (0.020) | 0.08 | 1290788 | 0.419 |
| avg. weight (\bar{Q}_{ij}) | 0.189 (0.011) | -0.058 (0.001) | 0.334 (0.037) | 0.087 (0.009) | -12.001 (0.031) | 0.05 | 1290788 | -0.537 |

Notes:

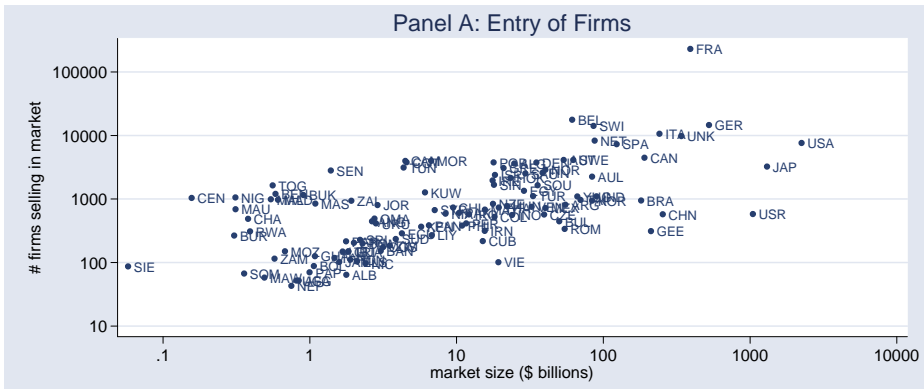
(a) Regression of (log) shipment value and its components from equations (7) and (8) on geographic variables. Dependent variables in left hand column. Coefficients in right-justified rows sum to coefficients in left justified rows.

(b) Standard errors in parentheses.

(c) ϵ_D is the elasticity of trade with respect to distance, evaluated at the sample mean distance of 523 miles.

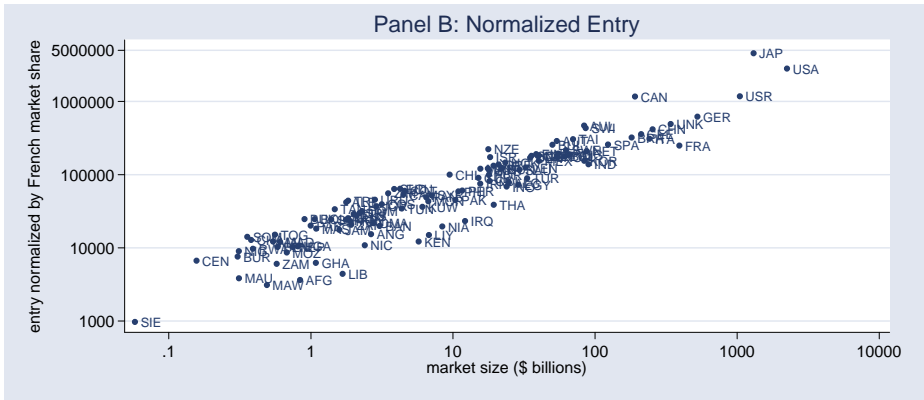
Eaton, Kortum and Kramarz (2009)

French Exporters: Extensive margin (N_{nF})



Eaton, Kortum and Kramarz (2009)

French Exporters: Extensive margin, normalized ($N_{nF}/(X_{nF}/X_n)$)



Helpman, Melitz and Rubenstein (QJE, 2008)

- What does the difference between intensive and extensive margins imply for the estimation of gravity equations?
 - Gravity equations are used to better understand trade theory (as in, for example, Evenett and Keller (2002)).
 - Gravity equations are often used as a reduced-form tool for measuring trade costs and the determinants of trade costs—we will see an entire lecture on estimating Trade Costs next week, and gravity equations will loom large.
- HMR (2008) started wave of thinking about gravity equation estimation in the presence of extensive/intensive margins.
 - They use aggregate international trade (so perhaps this paper doesn't really belong in a lecture on 'firm-level trade'!) to explore implications of a heterogeneous firm model for gravity equation estimation.
 - The Melitz (2003) model is simplified and used as a tool to understand, estimate, and correct for biases in gravity equation estimation.

HMR (2008): Zeros in Trade Data

- HMR start with the observation that there are lots of 'zeros' in international trade data, even when aggregated up to total bilateral exports.
 - Baldwin and Harrigan (2008) and Johnson (2008) look at this in a more disaggregated manner and find (unsurprisingly) far more zeros.
- Zeros are interesting.
- But zeros are also problematic.
 - Econometric: A typical analysis of trade flows is based on the gravity equation in logs, so observations with $X_{ij} = 0$ are censored out.
 - Theory: Models of the gravity equation (Armington, Krugman, Eaton-Kortum, Melitz with non-truncated Pareto) predict (for finite trade costs) that all countries trade all goods with all other countries (ie, no zeros).

HMR (2008)

The extent of zeros, even at the aggregate export level

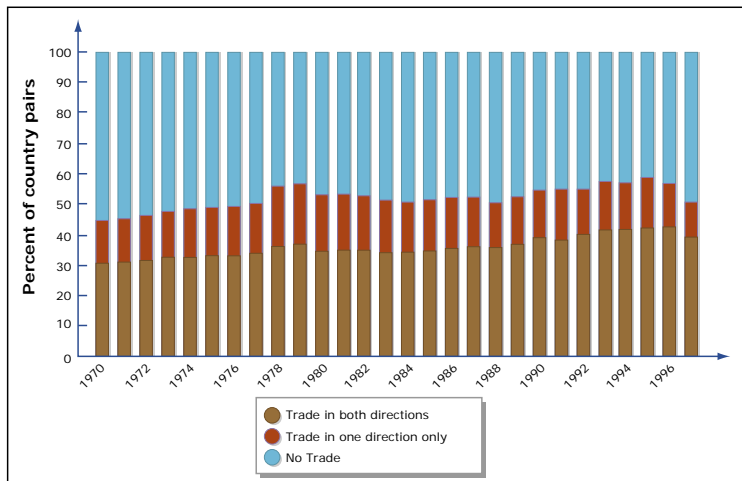


Image by MIT OpenCourseWare.

FIGURE I

Distribution of Country Pairs Based on Direction of Trade

Note. Constructed from 158 countries.

HMR (2008)

The growth of trade is not due to the death of zeros

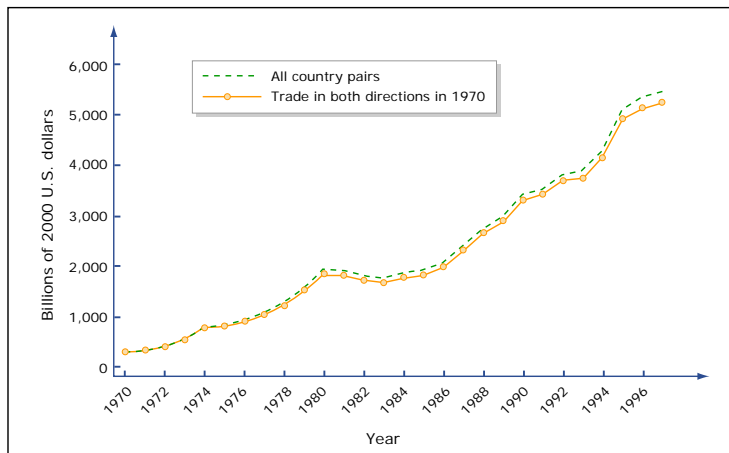


FIGURE II
Aggregate Volume of Exports of All Country Pairs and of Country Pairs That
Traded in Both Directions in 1970

A Gravity Model with Zeroes

- HMR work with a multi-country version of Melitz (2003)—similar to Chaney (2008).
- Set-up:
 - Monopolistic competition, CES preferences (ε), one factor of production (unit cost c_j), one sector.
 - Both variable (iceberg τ_{ij}) and fixed (f_{ij}) costs of exporting.
 - Heterogeneous firm-level productivities $1/a$ drawn from truncated Pareto, $G(a)$.
- Some firms in j sell in country i iff $a \leq a_{ij}$, where the cutoff productivity (a_{ij}) is defined by:

$$\kappa_1 \left(\frac{\tau_{ij} c_j a_{ij}}{P_i} \right)^{1-\varepsilon} Y_i = c_j f_{ij} \quad (1)$$

Trade Flows

- Bilateral exports from j to i are:

$$M_{ij} = \kappa_2 \left(\frac{c_j \tau_{ij}}{P_i} \right)^{1-\varepsilon} Y_i N_j V_{ij} \quad (2)$$

- Where $V_{ij} = \int_{a_L}^{a_{ij}} a^{1-\varepsilon} dG(a)$ if $a_{ij} \geq a_L$ and $V_{ij} = 0$ otherwise.
 - So this is an otherwise standard gravity equation, apart from the fact that M_{ij} can be zero.
 - And the important variable V_{ij} tells us where we should expect zeros.
- When $G(\cdot)$ is truncated Pareto (with parameter k), V_{ij} simplifies to be proportional to:

$$W_{ij} = \max \left\{ \left(\frac{a_{ij}}{a_L} \right)^{k+1-\varepsilon} - 1, 0 \right\} \quad (3)$$

The Gravity Equation

- Taking logs of the exporting equation, substituting in a trade costs parameterization ($\tau_{ij}^{\varepsilon-1} = D_{ij}^{\gamma} e^{-u_{ij}}$, where D is distance and $u_{ij} \sim N(0, \sigma_u^2)$) yields (where lower-case variables are in logs):

$$m_{ij} = \beta_0 + \alpha_i + \alpha_j - \gamma d_{ij} + w_{ij} + u_{ij} \quad (4)$$

- This is an unorthodox gravity equation because of the presence of w_{ij}
 - And of course, it is this term w_{ij} that accounts for selection (it is the log of W_{ij} which is just a transformation of V_{ij}).

Two Sources of Bias

- The HMR (2008) theory suggests (and solves) two sources of bias in the typical estimation of gravity equations.
- First: bias due to the presence of w_{ij} :
 - Above we saw that W_{ij} is a complicated function of a_{ij} which is itself a function of d_{ij} .
 - So estimation of the gravity equation without accounting for this will bias estimates of γ (OVB).
 - Intuitively, typical gravity equations assume that each firm ships the same amount. Here, W_{ij} corrects for the fact that productive firms ship more; and only productive firms penetrate distant markets.

Two Sources of Bias

- The HMR (2008) theory suggests (and solves) two sources of bias in the typical estimation of gravity equations.
- Second: A selection effect induced by only working with non-zero trade flows:
 - HMR's gravity equation, like those before it, can't be estimated on the observations for which $M_{ij} = 0$.
 - The HMR theory tells us that the existence of these 'zeros' is not as good as random with respect to d_{ij} , so econometrically this 'selection effect' needs to be corrected/controlled for.
 - Intuitively, the problem is that far away destinations are less likely to be profitable, so the sample of zeros is selected on the basis of d_{ij} .
 - This calls for a standard Heckman (1979) selection correction.

HMR (2008): Two-step Estimation

Two-step estimation to solve bias

1. Estimate probit for zero trade flow or not:
 - Include exporter and importer fixed effects, and d_{ij} .
 - Can proceed with just this, but then identification (in Step 2) is achieved purely off of the normality assumption.
 - To strengthen identification, need additional variable that enters Probit in step 1, but does not enter Step 2.
 - Theory says this should be a variable that affects the fixed cost of exporting, but not the variable cost.
 - HMR use Djankov et al (QJE, 2002)'s 'entry regulation' index. Also try 'common religion dummy.'
2. Estimate gravity equation on positive trade flows:
 - Include inverse Mills ratio (standard Heckman trick) to control for selection problem.
 - Also include empirical proxy for w_{ij} based on estimate of entry equation in Step 1.

HMR (2008): Results (traditional gravity estimation)

Benchmark Gravity and Selection into Trading Relationship

| Variables | 1986 | | 1980's | | | |
|--------------------|---------------------|----------------------|---------------------|----------------------|---------------------|----------------------|
| | m_{ij} | (Porbit) T_{ij} | m_{ij} | (Porbit) T_{ij} | m_{ij} | (Porbit) T_{ij} |
| Distance | -1.176** (0.031) | -0.263** (0.012) | -1.201** (0.024) | -0.246** (0.008) | -1.200** (0.024) | -0.246** (0.008) |
| Land border | 0.458** (0.147) | -0.148** (0.047) | 0.366** (0.131) | -0.146** (0.032) | 0.364** (0.131) | -0.146** (0.032) |
| Island | -0.391** (0.121) | -0.136** (0.032) | -0.381** (0.096) | -0.140** (0.022) | -0.378** (0.096) | -0.140** (0.022) |
| Landlock | -0.561** (0.188) | -0.072 (0.045) | -0.582** (0.148) | -0.087** (0.028) | -0.581** (0.147) | -0.087** (0.028) |
| Legal | 0.486** (0.050) | 0.038** (0.014) | 0.406** (0.040) | 0.029** (0.009) | 0.407** (0.040) | 0.028** (0.009) |
| Language | 1.176** (0.061) | 0.113** (0.016) | 0.207** (0.047) | 0.109** (0.011) | 0.203** (0.047) | 0.108** (0.011) |
| Colonial ties | 1.299** (0.120) | 0.128 (0.117) | 1.321** (0.110) | 0.114 (0.082) | 1.326** (0.110) | 0.116 (0.082) |
| Currency union | 1.364** (0.255) | 0.190** (0.052) | 1.395** (0.187) | 0.206** (0.026) | 1.409** (0.187) | 0.206** (0.026) |
| FTA | 0.759** (0.222) | 0.494** (0.020) | 0.996** (0.213) | 0.497** (0.018) | 0.976** (0.214) | 0.495** (0.018) |
| Religion | 0.102 (0.096) | 0.104** (0.025) | -0.018 (0.076) | 0.099** (0.016) | -0.038 (0.077) | 0.098** (0.016) |
| WTO (none) | | | | | -0.068 (0.058) | -0.056** (0.013) |
| WTO (both) | | | | | 0.303** (0.042) | 0.093** (0.013) |
| Observations R^2 | 11,146 0.709 | 24,649 0.587 | 110,697 0.682 | 248,060 0.551 | 110,697 0.682 | 248,060 0.551 |

Notes. Exporter, importer, and year fixed effects. Marginal effects at sample means and pseudo R^2 reported for Probit. Robust standard errors (clustering by country pair).

+ Significant at 10%

* Significant at 5%

** Significant at 1%

HMR (2008): Results (gravity estimation with correction)

| Baseline Results | | | | | | |
|--------------------------------------|----------------------|---------------------|--------------------|---------------------|---------------------|---------------------|
| Variables | 1986 reduced sample | | | | | |
| | (Probit) T_{ij} | m_{ij} | | | | |
| | | Benchmark | NLS | Polynomial | Indicator variables | |
| | | | | 50 bins | 100 bins | |
| Distance | -0.213** (0.016) | -1.167** (0.040) | -0.813 (0.049) | -0.847** (0.052) | -0.755** (0.070) | -0.789** (0.088) |
| Land border | -0.087 (0.072) | 0.627** (0.165) | 0.871 (0.170) | 0.845** (0.166) | 0.892** (0.170) | 0.863** (0.170) |
| Island | -0.173* (0.078) | -0.553* (0.269) | -0.203 (0.290) | -0.218 (0.258) | -0.161 (0.259) | -0.197 (0.258) |
| Landlock | -0.053 (0.050) | -0.432* (0.189) | -0.347* (0.175) | -0.362+ (0.187) | -0.352+ (0.187) | -0.353+ (0.187) |
| Legal | 0.049** (0.019) | 0.535** (0.064) | 0.431** (0.065) | 0.434** (0.064) | 0.407** (0.065) | 0.418** (0.065) |
| Language | 0.101** (0.021) | 0.147+ (0.075) | -0.030 (0.087) | -0.017 (0.077) | -0.061 (0.079) | -0.036 (0.083) |
| Colonial ties | -0.009 (0.130) | 0.909** (0.158) | 0.847** (0.257) | 0.848** (0.148) | 0.853** (0.152) | 0.838** (0.153) |
| Currency union | 0.216** (0.038) | 1.534** (0.334) | 1.077** (0.360) | 1.150** (0.333) | 1.045** (0.337) | 1.107** (0.346) |
| FTA | 0.343** (0.009) | 0.976** (0.247) | 0.124 (0.227) | 0.241 (0.197) | -0.141 (0.250) | 0.065 (0.348) |
| Religion | 0.141** (0.034) | 0.281* (0.120) | 0.120 (0.136) | 0.139 (0.120) | 0.073 (0.124) | 0.100 (0.128) |
| Regulation costs | -0.108** (0.036) | -1.146 (0.100) | — | — | — | — |
| R costs (days & proc) | -0.061* (0.031) | -0.216+ (0.124) | — | — | — | — |
| δ (from $\hat{\alpha}_{ij}$) | — | — | 0.840** (0.043) | — | — | — |
| $\hat{\eta}_{ij}^+$ | — | — | 0.240* (0.099) | 0.882** (0.209) | — | — |
| $\hat{\zeta}_{ij}^+$ | — | — | — | 3.261** (0.540) | — | — |
| $\hat{\zeta}_{ij}^{*2}$ | — | — | — | -0.712** (0.170) | — | — |
| $\hat{\zeta}_{ij}^{*3}$ | — | — | — | 0.060** (0.017) | — | — |
| Observations R^2 | 12,198 0.573 | 6,602 0.693 | 6,602 | 6,602 0.701 | 6,602 0.704 | 6,602 0.706 |

Notes: Exporter and importer fixed effects. Marginal effects at sample means and pseudo R^2 reported for Probit. Regulation costs are excluded variables in all second stage specifications. Bootstrapped standard errors for NLS; robust standard errors (clustering by country pair) elsewhere.

+ Significant at 10%.
* Significant at 5%.
** Significant at 1%.

- CK (2010) conduct a similar exercise to HMR (2008), but with French firm-level data.
 - This is attractive—after all, the main point that HMR (2008) is making is that firm-level realities matter for aggregate flows.
 - Hence looking at the firm-level adds certainty.
- CK's firm data has exports to foreign countries in it (CK focus only on adjacent countries: Belgium, Switzerland, Germany, Spain and Italy).

CK (2010): Identification

- But interestingly, CK also know where the firm is in France.
- So they try to separately identify the effects of variable and fixed trade costs by assuming:
 - Variable trade costs are proportional to distance. Since each firm is a different distance from, say, Belgium, there is cross-firm variation here.
 - Fixed trade costs are homogeneous across France for a given export destination. (It costs just as much to figure out how to sell to the Swiss whether your French firm is based in Geneva or Normandy).

- The model is deliberately close to Chaney (2008).
- In Chaney the variable trade cost (ie distance here, if we assume $\tau_{ij} = \theta D_{ij}^{\delta}$) elasticities of interest are:
 - Extensive: $\varepsilon_{D_{ij}}^{EXT_j} = -\delta [\gamma - (\sigma - 1)]$. CK estimate this by regressing firm-level entry (ie a Probit) on firm-level distance D_{ij} and a firm fixed effect. This is analogous to HMR's first stage.
 - Intensive (a la Krugman): $\varepsilon_{D_{ij}}^{INT_j} = -\delta(\sigma - 1)$. CK estimate this by regressing firm-level exports on firm-level distance D_{ij} and a firm fixed effect. This is analogous to HMR's second stage.

CK (2010): The model and estimation I

- Here γ is the Pareto parameter governing firm heterogeneity.
- The above two equations (HMR's first and second stage) don't separately identify δ , σ and γ .
 - So to identify the model, CK bring in another equation which is the firm size distribution.
 - In the Chaney (2008) model this will behave as:
$$X_i(\omega) = \lambda(1/a_i(\omega))^{-[\gamma-(\sigma-1)]}$$
, where ω indexes the firm.
 - With an Olley and Pakes (1996) TFP estimate of $1/a_i(\omega)$, CK estimate $[\gamma - (\sigma - 1)]$ and hence identify the entire system of 3 unknowns.

CK (2010): Results (each industry separately)

**The Structural Parameters of the Gravity Equation
(Firm-level Estimations)**

| Code | Industry | $P[\text{Export} > 0]$ $-\delta\gamma$ | Export value $-\delta(\sigma-1)$ | Pareto [#] $-\gamma-(\sigma-1)$ | γ | σ | δ |
|------|------------------------------|---|-------------------------------------|---|----------|----------|----------|
| 10 | Iron and steel | -5.51* | -1.71* | -1.36 | 1.98 | 1.62 | 2.78 |
| 11 | Steel processing | -1.5* | -0.99* | -1.74 | 5.1 | 4.36 | 0.29 |
| 13 | Metallurgy | -2.14* | -0.73* | -1.85 | 2.82 | 1.97 | 0.76 |
| 14 | Minerals | -2.98* | -0.91* | -2.86 | 4.11 | 2.25 | 0.72 |
| 15 | Ceramic and building mat. | -2.63* | -0.76* | -1.97 | 2.76 | 1.79 | 0.95 |
| 16 | Glass | -2.33* | -0.58* | -2.13 | 2.84 | 1.7 | 0.82 |
| 17 | Chemicals | -1.81* | -0.76* | -1.09 | 1.89 | 1.8 | 0.95 |
| 18 | Speciality chemicals | -0.97* | 0.34* | -1.39 | 2.13 | 1.74 | 0.46 |
| 19 | Pharmaceuticals | -1.19* | -0.14 | -1.4 | — | — | — |
| 20 | Foundry | -1.72* | -0.85* | -2.37 | 4.68 | 3.31 | 0.37 |
| 21 | Metal work | -1.19* | -0.36* | -2.43 | 3.48 | 2.05 | 0.34 |
| 22 | Agricultural machines | -2.06* | -0.57* | -2.39 | 3.31 | 1.92 | 0.62 |
| 23 | Machine tools | -1.29* | -0.48* | -2.47 | 3.92 | 2.45 | 0.33 |
| 24 | Industrial equipment | -1.25* | -0.48* | -1.97 | 3.21 | 2.24 | 0.39 |
| 25 | Mining / civil engnrng eqpmt | -1.37* | -0.46* | -1.9 | 2.86 | 1.96 | 0.48 |
| 27 | Office equipment | -0.52* | -1.02 | -1.57 | — | — | — |
| 28 | Electrical equipment | -0.8* | -0.14 | -2.34 | — | — | — |
| 29 | Electronical equipment | -0.77* | -0.24* | -1.63 | 2.34 | 1.71 | 0.33 |
| 30 | Domestic equipment | -0.94* | -0.14* | -2.13 | 2.51 | 1.37 | 0.38 |
| 31 | Transport equipment | -1.4* | -0.55* | -2.23 | 3.69 | 2.46 | 0.38 |
| 32 | Ship building | -3.69* | -2.67* | -1.52 | 5.53 | 5.01 | 0.67 |
| 33 | Aeronautical building | -0.78* | -0.13 | -3.27 | — | — | — |
| 34 | Precision instruments | -1.07* | 0.08 | -1.63 | — | — | — |
| 44 | Textile | -1.17* | -0.3* | -1.37 | 1.84 | 1.47 | 0.64 |
| 45 | Leather products | -1.24* | -0.44* | -1.63 | 2.53 | 1.9 | 0.49 |
| 46 | Shoe industry | -0.42* | -0.29* | -2.3 | 7.31 | 6.01 | 0.06 |
| 47 | Garment industry | -0.33* | 0.13 | -1.04 | — | — | — |
| 48 | Mechanical woodwork | -2.14* | -0.2* | -1.5 | 1.65 | 1.15 | 1.29 |
| 49 | Furniture | -1.43* | -0.37* | -2.25 | 3.04 | 1.79 | 0.47 |
| 50 | Paper & Cardboard | -1.45* | 0.76* | -1.76 | 3.71 | 2.95 | 0.39 |
| 51 | Printing and editing | -1.4* | 0.7* | -1.24 | 2.46 | 2.22 | 0.57 |
| 52 | Rubber | -1.26* | 0.8* | -2.52 | 6.93 | 5.41 | 0.18 |
| 53 | Plastic processing | -1.24* | 0.51* | -1.6 | 2.7 | 2.11 | 0.46 |
| 54 | Miscellaneous | -0.91* | -0.33* | -1.22 | 1.92 | 1.7 | 0.47 |
| | Tread weighted mean | -1.41 | -0.53 | -1.86 | 3.09 | 2.25 | 0.58 |

*, ** and *** denote significance at the 1%, 5% and 10% level respectively. #: All coefficients in this column are significant at the 1% level. Estimations include the contiguity variable.

CK (2010): Results (do the parameters make sense?)

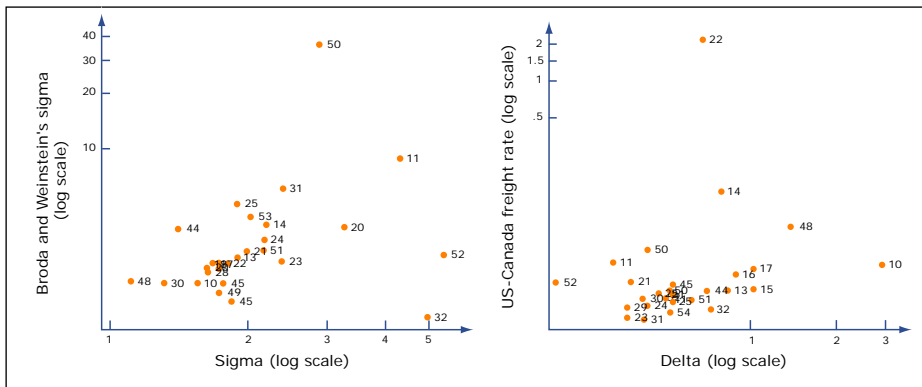


Image by MIT OpenCourseWare.

Figure 3: Comparison of our results for σ and δ with those of Broda and Weinstein (2003)

CK (2010): Results (what do the parameters imply about margins?)

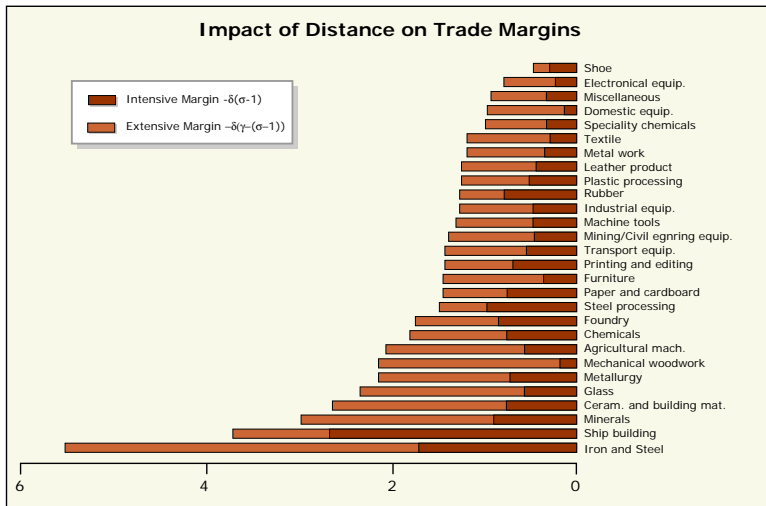
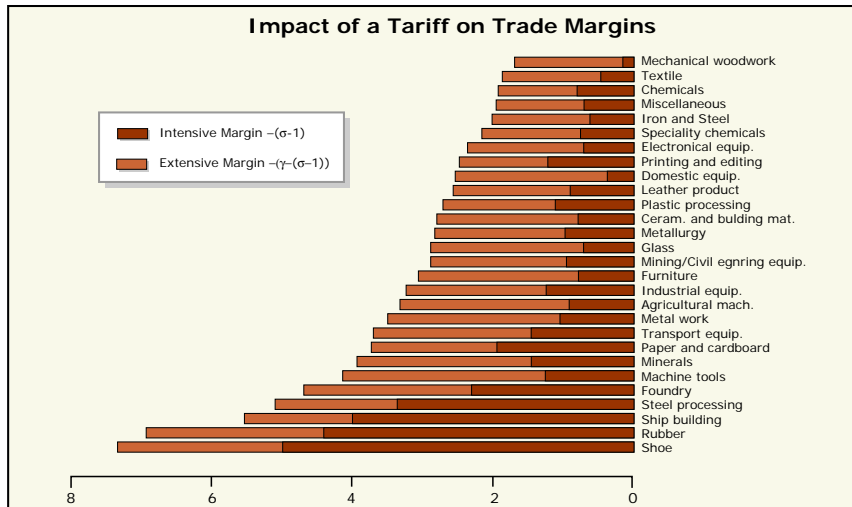


Image by MIT OpenCourseWare.

Figure 4: The estimated impact of trade barriers and distance on trade margins, by industry

CK (2010): Results (what do the parameters imply about margins?)

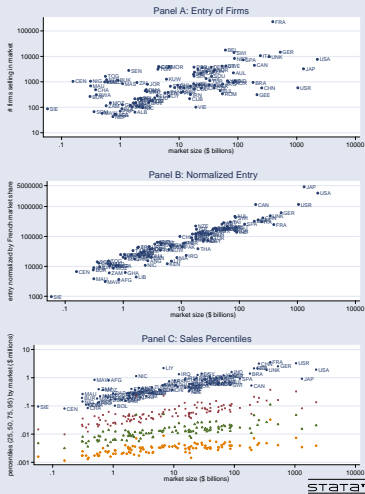


- EKK (2009) construct a Melitz-like model in order to try to capture the key features of French firms' exporting behavior:
 - Whether to export (simple extensive margin).
 - Which countries to export to (country-wise extensive margins).
 - How much to export to each country (intensive margin).
- They uncover some striking regularities in the firm-wise sales data in (multiple) foreign markets.
 - These 'power law' like relationships occur all over the place (Gabaix (ARE survey, 2009)).
 - Most famously, they occur for domestic sales within one market.
 - In that sense, perhaps it's not surprising that they also occur market by market abroad. (At the heart of power laws is the property of scale invariance.)

EKK (2009): Stylised Fact 1: Market Entry (averages across countries)

'Normalization': $N_{nF}/(X_{nF}/X_n)$

Figure 1: Entry and Sales by Market Size



EKK (2009): Stylised Fact 1: Market Entry (averages across countries)

All exporters export to at least one of these 7 places. But it's not a strict hierarchy as one would see in Melitz (2003).

French Firms Exporting to the Seven Most Popular Destinations

| Country | Number of exporters | Fraction of exporters |
|------------------------|---------------------|-----------------------|
| Belgium* (BE) | 17,699 | 0.520 |
| Germany (DE) | 14,579 | 0.428 |
| Switzerland (CH) | 14,173 | 0.416 |
| Italy (IT) | 10,643 | 0.313 |
| United Kingdom (UK) | 9,752 | 0.287 |
| Netherlands (NL) | 8,294 | 0.244 |
| United States (US) | 7,608 | 0.224 |
| Total Exporters | 34,035 | |

* Belgium includes Luxembourg

EKK (2009): Stylised Fact 1: Market Entry (averages across countries)

For 27% of exporters, a strict hierarchy is observed over these 7 destinations. Within these firms, foreign market entry is not independent.

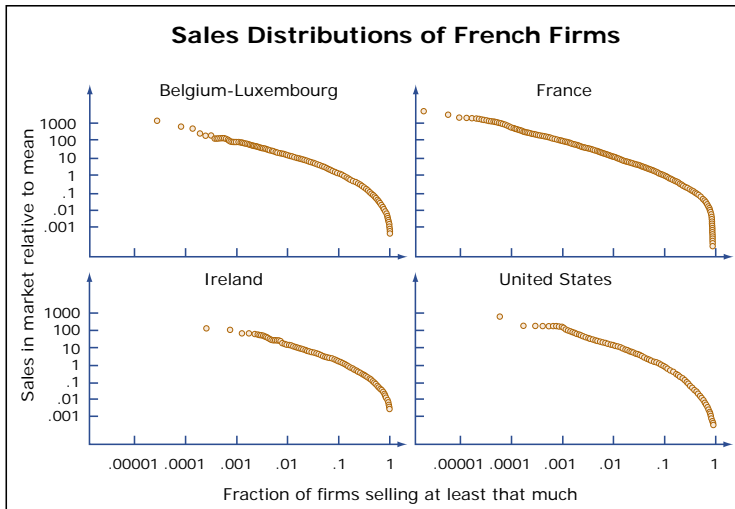
French Firms Selling to Strings of Top Seven Countries

| Export string | Number of French exporters | | |
|----------------------|----------------------------|--------------------|--------------|
| | Data | Under independence | Model |
| BE* | 3,988 | 1,700 | 4,417 |
| BE-DE | 863 | 1,274 | 912 |
| BE-DE-CH | 579 | 909 | 402 |
| BE-DE-CH-IT | 330 | 414 | 275 |
| BE-DE-CH-IT-UK | 313 | 166 | 297 |
| BE-DE-CH-IT-UK-NL | 781 | 54 | 505 |
| BE-DE-CH-IT-UK-NL-US | 2,406 | 15 | 2,840 |
| Total | 9,260 | 4,532 | 9,648 |

* The string "BE" means selling to Belgium but no other among the top 7, "BE-DE" means selling to Belgium and Germany but no other, etc.

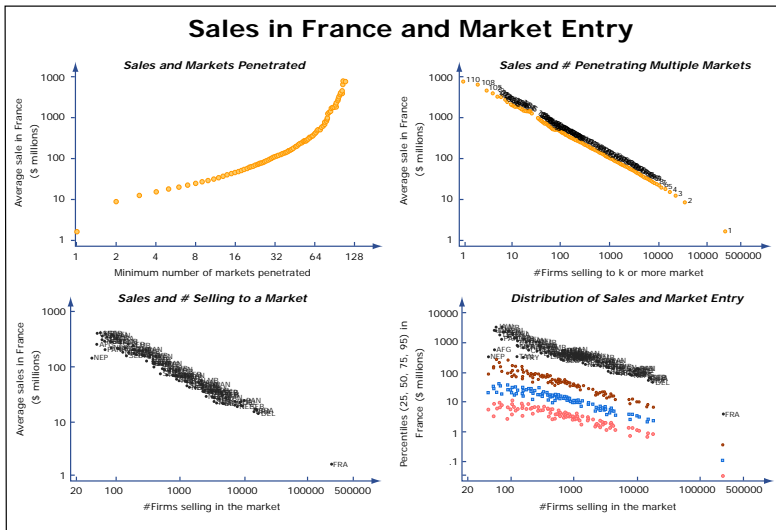
EKK (2009): Stylised Fact 2: Sales Distributions (across all firms)

Surprisingly similar shape (with 'mean' shift) in each destination market (including home). Power laws (at least in upper tails).



EKK (2009): Stylised Fact 3: Export Participation and Size in France

Big firms at home are multi-destination exporters.



EKK (2009): Model

- The above relationships fit the Melitz (2003) model (with $G(\cdot)$ being Pareto) in some regards, but not all.
- EKK (2009) therefore add some features to Melitz (2003) in order to bring this model closer to the data.
- Most of these will take the flavor of ‘firm-specific shocks/noise’.
 - The shocks smooths things out, allows for unobserved heterogeneity, and answer the structural econometrician’s question of “where does your regression’s error term come from?” .

EKK (2009) Model

- Shocks:
 - Firm (ie j)-specific productivity draws (in country i): $z_i(j)$. This is Pareto with parameter θ .
 - Firm-specific demand draw $\alpha_n(j)$. The demand they face in market n is thus: $X_n(j) = \alpha_n(j) f X_n \left(\frac{p}{P_n} \right)^{-(\sigma-1)}$, where f will be defined shortly.
 - Firm-specific fixed entry costs $E_{ni}(j) = \varepsilon_n(j) E_{ni} M(f)$, where $\varepsilon_n(j)$ is the firm-specific 'fixed exporting cost shock', E_{ni} is the fixed exporting term that appears in Melitz (2003) or HMR (2008) (ie constant across firms). And $M(f) = \frac{1-(1-f)^{1-1/\lambda}}{1-1/\lambda}$, which, following Arkolakis (2011), is a micro-founded 'marketing' function that captures how much firms have to pay to 'access' f consumers (this is a choice variable).
 - EKK assume that $g(\alpha, \varepsilon)$ can take any form, but it needs to be the same across countries n , iid across firms, and within firms independent from the Pareto distribution of z .

EKK (2009) Model: Entry

- The entry condition is similar to Melitz (2003). Enter if cost $c_{ni}(j) = \frac{w_i \tau_{ij}}{z_i(j)}$ satisfies:

$$c \leq c_{ni}(\eta) \equiv \left(\frac{\eta X_n}{\sigma E_{ni}} \right)^{1/(\sigma-1)} \frac{P_n}{m} \quad (5)$$

- Here $\eta_n(j) \equiv \frac{\alpha_n(j)}{\varepsilon_n(j)}$.
 - And X_n is total sales in n , P_n is the price index in n , and m is the (constant) markup.
- Integrating this over the distribution $g(\eta)$ we know how much entry (measure of firms) there is:

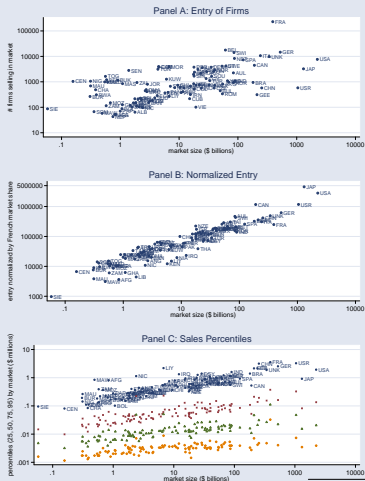
$$J_{ni} = \frac{\kappa_2 \pi_{ni} X_n}{\kappa_1 \sigma E_{ni}} \quad (6)$$

- This therefore agrees well with Fact 1 (normalized entry is linear in X_n).

EKK (2009): Stylised Fact 1: Market Entry (averages across countries)

'Normalization': $N_{nF}/(X_{nF}/X_n)$

Figure 1: Entry and Sales by Market Size



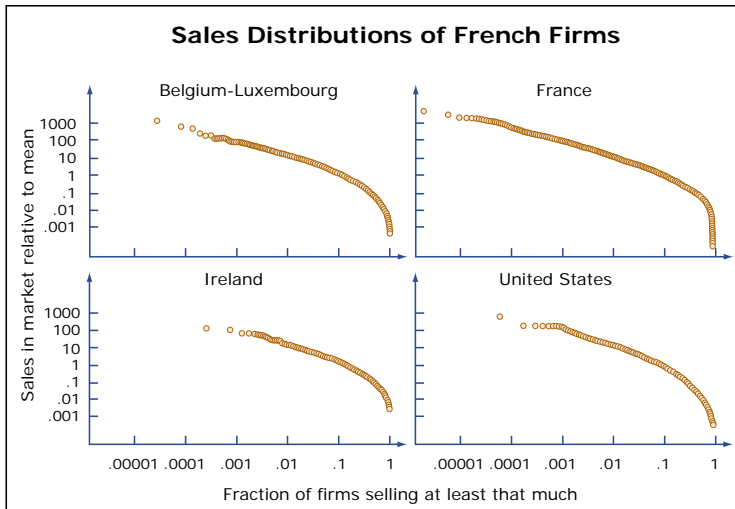
- The firm sales (conditional on entry) condition is similar to Arkolakis (2011):

$$X_{ni}(j) = \varepsilon \left[1 - \left(\frac{c}{c_{ni}(\eta)} \right)^{\lambda(\sigma-1)} \right] \left(\frac{c}{c_{ni}(\eta)} \right)^{-(\sigma-1)} \sigma E_{ni}. \quad (7)$$

- There is more work to be done, but one can already see that this will look a lot like a Pareto distribution (c is Pareto, so c to any power is also Pareto) in each market (as in Figure 2).
- But the $\left[1 - \left(\frac{c}{c_{ni}(\eta)} \right)^{\lambda(\sigma-1)} \right]$ will cause the sales distribution to deviate from Pareto in the lower tail (also as in Figure 2).

EKK (2009): Stylised Fact 2: Sales Distributions (across all firms)

Surprisingly similar shape (with 'mean' shift) in each destination market (including home). Power laws (at least in upper tails).



EKK (2009) Model: Sales in France Conditional on Foreign Entry

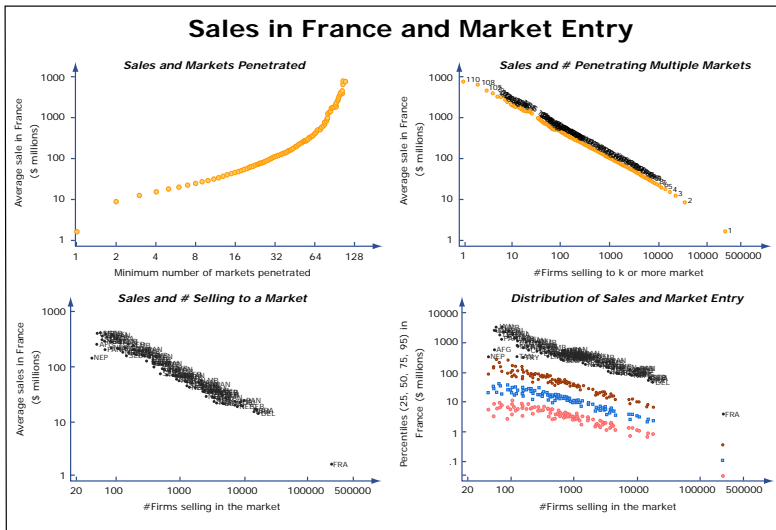
- The amount of sales in France conditional on entering market n can be shown to be:

$$X_{FF}(j)|_n = \frac{\alpha_F(j)}{\eta_n(j)} \left[1 - v_{nF}(j)^{\lambda/\tilde{\theta}} \left(\frac{N_{nF}}{N_{FF}} \right)^{\lambda/\tilde{\theta}} \left(\frac{\eta_n(j)}{\eta_F(j)} \right)^\lambda \right] \\ \times v_{nF}(j)^{-1/\tilde{\theta}} \left(\frac{N_{nF}}{N_{FF}} \right)^{-1/\tilde{\theta}} \frac{\kappa_2}{\kappa_1} X_{FF}.$$

- Since N_{nF}/N_{FF} is close to zero (everywhere but in France) the dependence of this on N_{nF} is Pareto with slope $-1/\tilde{\theta}$. As in Figure 3.

EKK (2009): Stylised Fact 3: Export Participation and Size in France

Big firms at home are multi-destination exporters.



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